

1) (a)  $f_{(x,y)}$  is obtained by dividing each entry in the table by the total 205 ①  
obtaining

$x \backslash y$	1	2	3	4
1	33/205	26/205	10/205	20/205
2	21/205	10/205	13/205	12/205
3	16/205	18/205	16/205	20/205

b) The conditional distributions are obtained by taking the original table and dividing each row by the row total.

$f_{y 1} (y   X=1)$	0.371	0.242	0.142	0.225
$f_{y 2} (y   X=2)$	0.375	0.179	0.232	0.214
$f_{y 3} (y   X=3)$	0.229	0.257	0.229	0.286

⇒) There is a relationship between  $X$  and  $Y$  because the conditional distributions are different.

Plotting the conditional distributions reveals no discernible pattern to this relationship.

⊃ We have that

$$\begin{aligned}
 P_{Y|X}(y|x) &= \frac{\#\{\omega : X(\omega) = x, Y(\omega) = y\}}{\#\{\omega : X(\omega) = x\}} \\
 &= \frac{\#\{\omega : X(\omega) = x, Y(\omega) = y\}}{\#\{\omega : X(\omega) = x\}} \bigg/ \frac{\#\{\omega : X(\omega) = x\}}{\#\{\omega\}} \\
 &= \frac{P_{X,Y}(x,y)}{P_X(x)} \\
 &\stackrel{\text{hypothesis}}{=} \frac{P_X(x) P_Y(y)}{P_X(x)} = P_Y(y)
 \end{aligned}$$

and therefore  $P_{Y|X}(y|x)$  doesn't change as we change  $x$  and there is no relationship.

⊃ Let  $g(y) = P_{Y|X}(y|x)$  as this doesn't depend on  $x$  by hypothesis. Since (as argued above)

$$\begin{aligned}
 P_{X,Y}(x,y) &= P_{Y|X}(y|x) P_X(x) \\
 &= g(y) P_X(x)
 \end{aligned}$$

$$\begin{aligned}
 \text{and so } P_Y(y) &= \sum_x P_{X,Y}(x,y) \\
 &= \sum_x g(y) P_X(x) = g(y) \left( \sum_x P_X(x) \right) \\
 &= g(y) \text{ since } \sum_x P_X(x) = \sum_x \frac{\#\{\omega : X(\omega) = x\}}{\#\{\omega\}} = 1
 \end{aligned}$$

$$\textcircled{3} \sum_{x \in (c_i, c_{i+1}]} f_x(x) = \sum_{x \in (c_i, c_{i+1}]} f_x(x) + \sum_{x \in (c_{i+1}, c_{i+2}]} f_x(x) + \dots + \sum_{x \in (c_{j-1}, c_j]} f_x(x)$$

$$\text{and } \sum_{x \in (c_i, c_{i+1}]} f_x(x) = \frac{\#\{\omega : X(\omega) \in (c_i, c_{i+1}]\}}{\#\{\Omega\}} \quad \text{Now}$$

$$\int_{c_i}^{c_{i+1}} f_x(x) dx = \int_{c_i}^{c_{i+1}} f_x(x) dx + \int_{c_{i+1}}^{c_{i+2}} f_x(x) dx + \dots + \int_{c_{j-1}}^{c_j} f_x(x) dx$$

$$\text{and } \int_{c_i}^{c_{i+1}} f_x(x) dx = \int_{c_i}^{c_{i+1}} \frac{\#\{\omega : X(\omega) \in (c_i, c_{i+1}]\}}{\#\{\Omega\}} dx$$

$$= \frac{\#\{\omega : X(\omega) \in (c_i, c_{i+1}]\}}{\#\{\Omega\}} \int_{c_i}^{c_{i+1}} dx$$

$$= \frac{\#\{\omega : X(\omega) \in (c_i, c_{i+1}]\}}{\#\{\Omega\}} \frac{(c_{i+1} - c_i)}{(c_{i+1} - c_i)} = \sum_{x \in (c_i, c_{i+1}]} f_x(x)$$

which establishes the result.

$\textcircled{4}$  (a) We have that

$$f_{Y|X}(y|x) = f_{(X,Y)}(x,y) / f_X(x)$$

and using  $X \sim N(\mu_1, \sigma_1^2)$  this implies

$$f_{Y|X}(y|x) \text{ using } \begin{pmatrix} \sigma_1^2 & \sigma_1 \sigma_2 \rho \\ \sigma_1 \sigma_2 \rho & \sigma_2^2 \end{pmatrix}^{-1} = \frac{1}{\sigma_1^2 \sigma_2^2 (1-\rho^2)} \begin{pmatrix} \sigma_2^2 & -\sigma_1 \sigma_2 \rho \\ -\sigma_1 \sigma_2 \rho & \sigma_1^2 \end{pmatrix}$$

$$= (2\pi)^{-1} (\sigma_1^2 \sigma_2^2 (1-\rho^2))^{-\frac{1}{2}} \exp\left\{ \frac{-1}{2(1-\rho^2)} \left[ \left(\frac{x-\mu_1}{\sigma_1}\right)^2 - 2\rho \left(\frac{x-\mu_1}{\sigma_1}\right) \left(\frac{y-\mu_2}{\sigma_2}\right) + \left(\frac{y-\mu_2}{\sigma_2}\right)^2 \right] \right\}$$

$$(2\pi)^{-\frac{1}{2}} (\sigma_1^2)^{-\frac{1}{2}} \exp\left\{ -\frac{1}{2} \left(\frac{x-\mu_1}{\sigma_1}\right)^2 \right\}$$

$$= (2\pi)^{-\frac{1}{2}} (\sigma_2^2(1-\rho^2))^{-\frac{1}{2}} x$$

$$e^{-\rho^2} \left\{ -\frac{1}{2(1-\rho^2)} \left[ \left(\frac{x-\mu_1}{\sigma_1}\right)^2 - 2\rho \left(\frac{x-\mu_1}{\sigma_1}\right) \left(\frac{y-\mu_2}{\sigma_2}\right) + \left(\frac{y-\mu_2}{\sigma_2}\right)^2 - (1-\rho^2) \left(\frac{x-\mu_1}{\sigma_1}\right)^2 \right] \right\}$$

and  $[ ] = \rho^2 \left(\frac{x-\mu_1}{\sigma_1}\right)^2 - 2\rho \left(\frac{x-\mu_1}{\sigma_1}\right) \left(\frac{y-\mu_2}{\sigma_2}\right) + \left(\frac{y-\mu_2}{\sigma_2}\right)^2$   
 $= \left[ \left(\frac{y-\mu_2}{\sigma_2}\right) - \rho \left(\frac{x-\mu_1}{\sigma_1}\right) \right]^2 = \frac{1}{\sigma_2^2} \left[ y - \mu_2 - \rho \frac{\sigma_2}{\sigma_1} (x - \mu_1) \right]^2$

$$\therefore Y|X=x \sim N \left( \mu_2 + \frac{\rho\sigma_2}{\sigma_1} (x-\mu_1), \sigma_2^2(1-\rho^2) \right)$$

(b) When  $\rho > 0$  then  $X$  and  $Y$  are correlated so they vary together. When we condition on  $X$ , namely fix  $X$ , then we are removing the joint variation from  $Y$  leaving only the variation in  $Y$  that has nothing to do with  $X$ .